



Model-Based Recursive Partitioning

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Overview

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Motivation: Trees

Breiman (2001, *Statistical Science*) distinguishes two cultures of statistical modeling.

- **Data models:** Stochastic models, typically parametric.
- **Algorithmic models:** Flexible models, data-generating process unknown.

Example: Recursive partitioning models dependent variable Y by “learning” a partition w.r.t explanatory variables Z_1, \dots, Z_l .

Key features:

- Predictive power in nonlinear regression relationships.
- Interpretability (enhanced by visualization), i.e., no “black box” methods.

Motivation: Leaves

Typically: Simple models for univariate Y , e.g., mean or proportion.

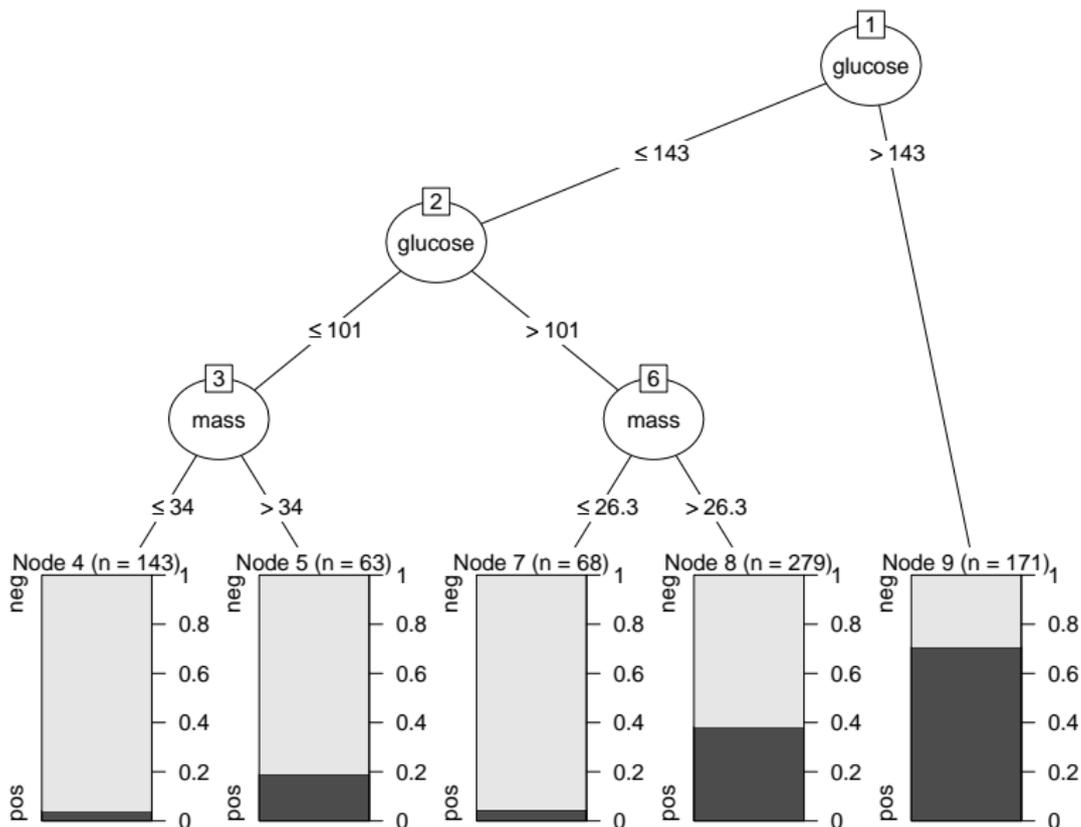
Examples: CART and C4.5 in statistical and machine learning, respectively.

Idea: More complex models for multivariate Y , e.g., multivariate normal model, regression models, etc.

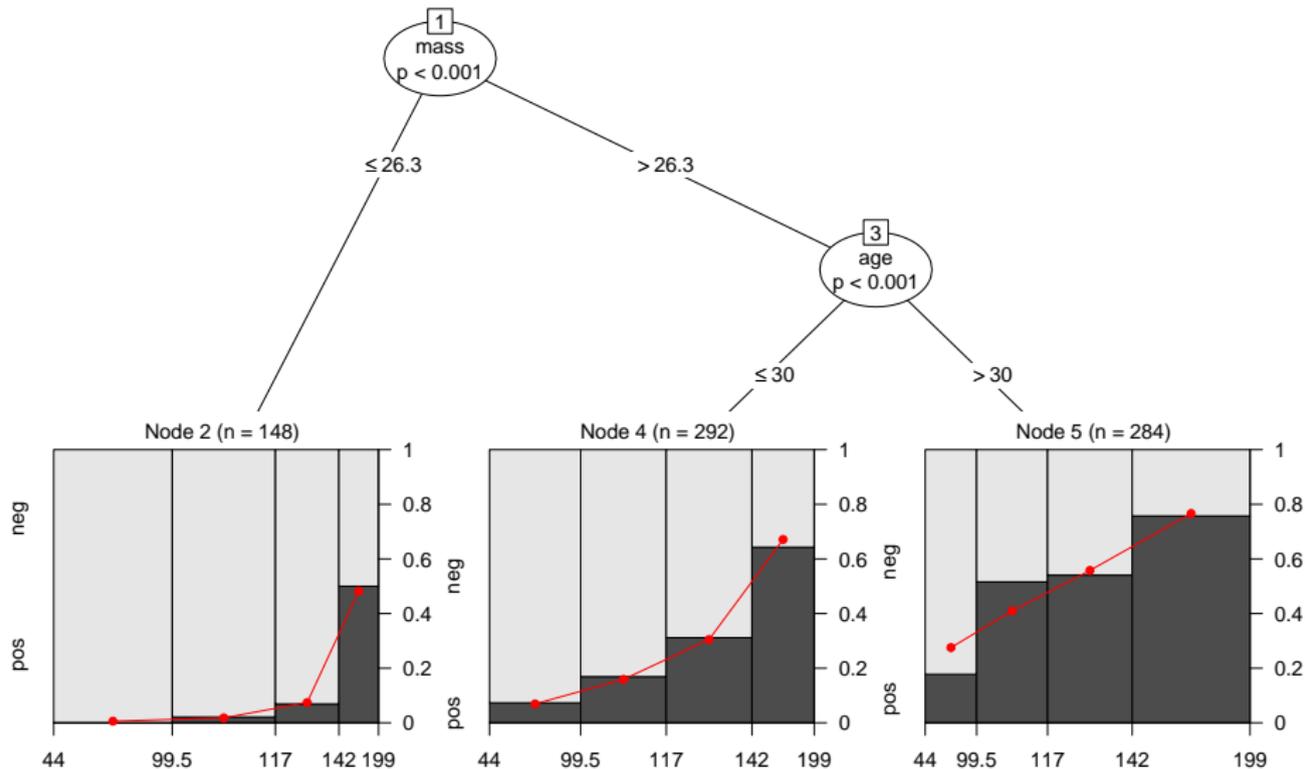
Here: Synthesis of parametric data models and algorithmic tree models.

Goal: Fitting local models by partitioning of the sample space.

Motivation: Leaves



Motivation: Leaves



Recursive partitioning

Base algorithm:

- 1 Fit model for Y .
- 2 Assess association of Y and each Z_j .
- 3 Split sample along the Z_{j^*} with strongest association: Choose breakpoint with highest improvement of the model fit.
- 4 Repeat steps 1–3 recursively in the sub-samples until some stopping criterion is met.

Here: Segmentation (3) of parametric models (1) with additive objective function using parameter instability tests (2) and associated statistical significance (4).

1. Model estimation

Models: $\mathcal{M}(Y, \theta)$ with (potentially) multivariate observations $Y \in \mathcal{Y}$ and k -dimensional parameter vector $\theta \in \Theta$.

Parameter estimation: $\hat{\theta}$ by optimization of objective function $\Psi(Y, \theta)$ for n observations Y_i ($i = 1, \dots, n$):

$$\hat{\theta} = \operatorname{argmin}_{\theta \in \Theta} \sum_{i=1}^n \Psi(Y_i, \theta).$$

Special cases: Maximum likelihood (ML), weighted and ordinary least squares (OLS and WLS), quasi-ML, and other M-estimators.

Central limit theorem: If there is a true parameter θ_0 and given certain weak regularity conditions, $\hat{\theta}$ is asymptotically normal with mean θ_0 and sandwich-type covariance.

1. Model estimation

Estimating function: $\hat{\theta}$ can also be defined in terms of

$$\sum_{i=1}^n \psi(Y_i, \hat{\theta}) = 0,$$

where $\psi(Y, \theta) = \partial\Psi(Y, \theta)/\partial\theta$.

Idea: In many situations, a single global model $\mathcal{M}(Y, \theta)$ that fits **all** n observations cannot be found. But it might be possible to find a partition w.r.t. the variables $Z = (Z_1, \dots, Z_l)$ so that a well-fitting model can be found locally in each cell of the partition.

Tool: Assess parameter instability w.r.t to partitioning variables $Z_j \in \mathcal{Z}_j$ ($j = 1, \dots, l$).

2. Tests for parameter instability

Generalized M-fluctuation tests capture instabilities in $\hat{\theta}$ for an ordering w.r.t Z_j .

Basis: Empirical fluctuation process of cumulative deviations w.r.t. to an ordering $\sigma(Z_{ij})$.

$$W_j(t, \hat{\theta}) = \hat{B}^{-1/2} n^{-1/2} \sum_{i=1}^{\lfloor nt \rfloor} \psi(Y_{\sigma(Z_{ij})}, \hat{\theta}) \quad (0 \leq t \leq 1)$$

Functional central limit theorem: Under parameter stability $W_j(\cdot) \xrightarrow{d} W^0(\cdot)$, where W^0 is a k -dimensional Brownian bridge.

2. Tests for parameter instability

Test statistics: Scalar functional $\lambda(W_j)$ that captures deviations from zero.

Null distribution: Asymptotic distribution of $\lambda(W^0)$.

Special cases: Class of test encompasses many well-known tests for different classes of models. Certain functionals λ are particularly intuitive for numeric and categorical Z_j , respectively.

Advantage: Model $\mathcal{M}(Y, \hat{\theta})$ just has to be estimated once. Empirical estimating functions $\psi(Y_i, \hat{\theta})$ just have to be re-ordered and aggregated for each Z_j .

2. Tests for parameter instability

Splitting numeric variables: Assess instability using sup LM statistics.

$$\lambda_{\text{sup}LM}(W_j) = \max_{i=\underline{i}, \dots, \bar{i}} \left(\frac{i}{n} \cdot \frac{n-i}{n} \right)^{-1} \left\| W_j \left(\frac{i}{n} \right) \right\|_2^2.$$

Interpretation: Maximization of single shift LM statistics for all conceivable breakpoints in $[\underline{i}, \bar{i}]$.

Limiting distribution: Supremum of a squared, k -dimensional tied-down Bessel process.

2. Tests for parameter instability

Splitting categorical variables: Assess instability using χ^2 statistics.

$$\lambda_{\chi^2}(W_j) = \sum_{c=1}^C \frac{n}{|I_c|} \left\| \Delta_{I_c} W_j \left(\frac{i}{n} \right) \right\|_2^2$$

Feature: Invariant for re-ordering of the C categories and the observations within each category.

Interpretation: Captures instability for split-up into C categories.

Limiting distribution: χ^2 with $k \cdot (C - 1)$ degrees of freedom.

3. Segmentation

Goal: Split model into $b = 1, \dots, B$ segments along the partitioning variable Z_j associated with the highest parameter instability. Local optimization of

$$\sum_b \sum_{i \in I_b} \Psi(Y_i, \theta_b).$$

$B = 2$: Exhaustive search of order $O(n)$.

$B > 2$: Exhaustive search is of order $O(n^{B-1})$, but can be replaced by dynamic programming of order $O(n^2)$. Different methods (e.g., information criteria) can choose B adaptively.

Here: Binary partitioning.

4. Pruning

Pruning: Avoid overfitting.

Pre-pruning: Internal stopping criterion. Stop splitting when there is no significant parameter instability.

Post-pruning: Grow large tree and prune splits that do not improve the model fit (e.g., via cross-validation or information criteria).

Here: Pre-pruning based on Bonferroni-corrected p values of the fluctuation tests.

Costly journals

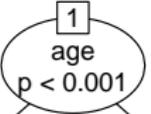
Task: Price elasticity of demand for economics journals.

Source: Bergstrom (2001, *Journal of Economic Perspectives*) “Free Labor for Costly Journals?”, used in Stock & Watson (2007), *Introduction to Econometrics*.

Model: Linear regression via OLS.

- Demand: Number of US library subscriptions.
- Price: Average price per citation.
- Log-log-specification: Demand explained by price.
- Further variables without obvious relationship: Age (in years), number of characters per page, society (factor).

Costly journals

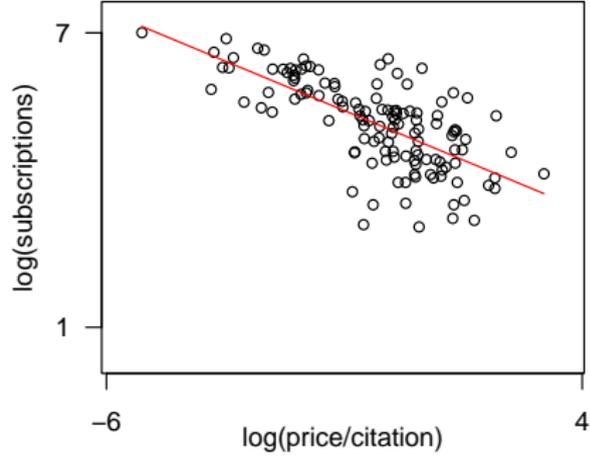
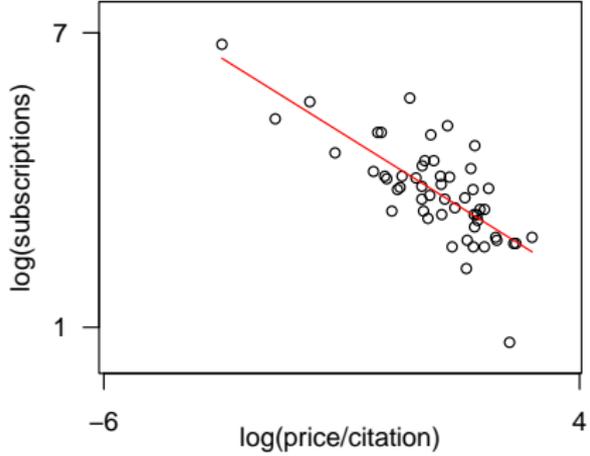


≤ 18

> 18

Node 2 (n = 53)

Node 3 (n = 127)



Costly journals

Recursive partitioning:

	Regressors		Partitioning variables				
	(Const.)	log(Pr./Cit.)	Price	Cit.	Age	Chars	Society
1	4.766	-0.533	3.280	5.261	42.198	7.436	6.562
	< 0.001	< 0.001	0.660	0.988	< 0.001	0.830	0.922
2	4.353	-0.605	0.650	3.726	5.613	1.751	3.342
	< 0.001	< 0.001	0.998	0.998	0.935	1.000	1.000
3	5.011	-0.403	0.608	6.839	5.987	2.782	3.370
	< 0.001	< 0.001	0.999	0.894	0.960	1.000	1.000

(Wald tests for regressors, parameter instability tests for partitioning variables.)

Pima Indians diabetes

Task: Classification of diabetes in Pima Indian women.

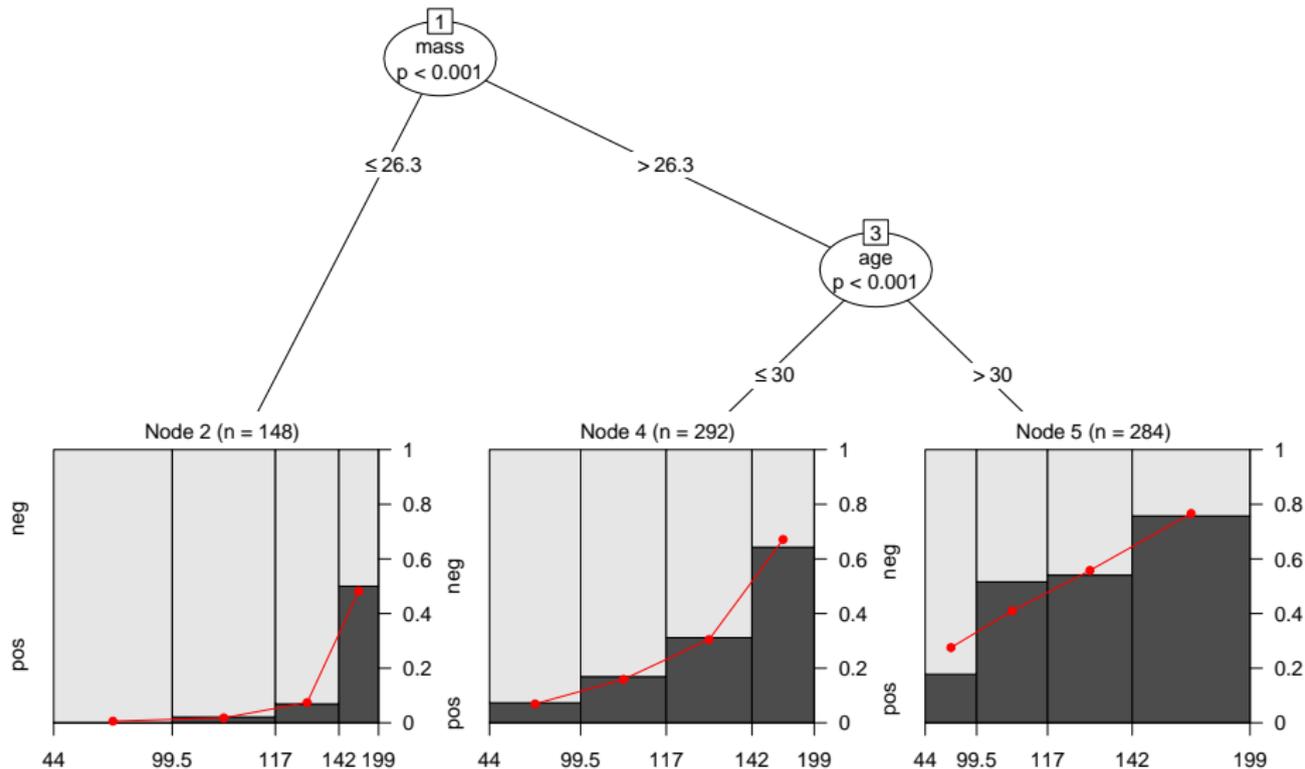
Source: Asuncion & Newman (2007), UCI Repository of Machine Learning Databases.

<http://www.ics.uci.edu/~mllearn/MLRepository.html>.

Model: Logistic regression via ML.

- Response: Test result for diabetes (positive/negative).
- Regressor: Plasma glucose concentration.
- Partitioning variables: Body mass index, age, number of pregnancies, blood pressure, diabetes pedigree function.

Pima Indians diabetes



Pima Indians diabetes

Recursive partitioning:

	(Constant)	Glucose conc.
2	-10.999	0.065
4	-6.573	0.045
5	-3.319	0.027

Beauty and teaching evaluation

Task: Correlation of beauty and teaching evaluations for professors.

Source: Hamermesh & Parker (2005, *Economics of Education Review*). “Beauty in the Classroom: Instructors’ Pulchritude and Putative Pedagogical Productivity.”

Model: Linear regression via WLS.

- Response: Average teaching evaluation per course (on scale 1–5).
- Explanatory variables: Standardized measure of beauty and factors gender, minority, tenure, etc.
- Weights: Number of students per course.

Beauty and teaching evaluation

	All	Men	Women
(Constant)	4.216	4.101	4.027
Beauty	0.283	0.383	0.133
Gender (= w)	-0.213		
Minority	-0.327	-0.014	-0.279
Native speaker	-0.217	-0.388	-0.288
Tenure track	-0.132	-0.053	-0.064
Lower division	-0.050	0.004	-0.244
R^2	0.271	0.316	

(Remark: Only courses with more than a single credit point.)

Beauty and teaching evaluation

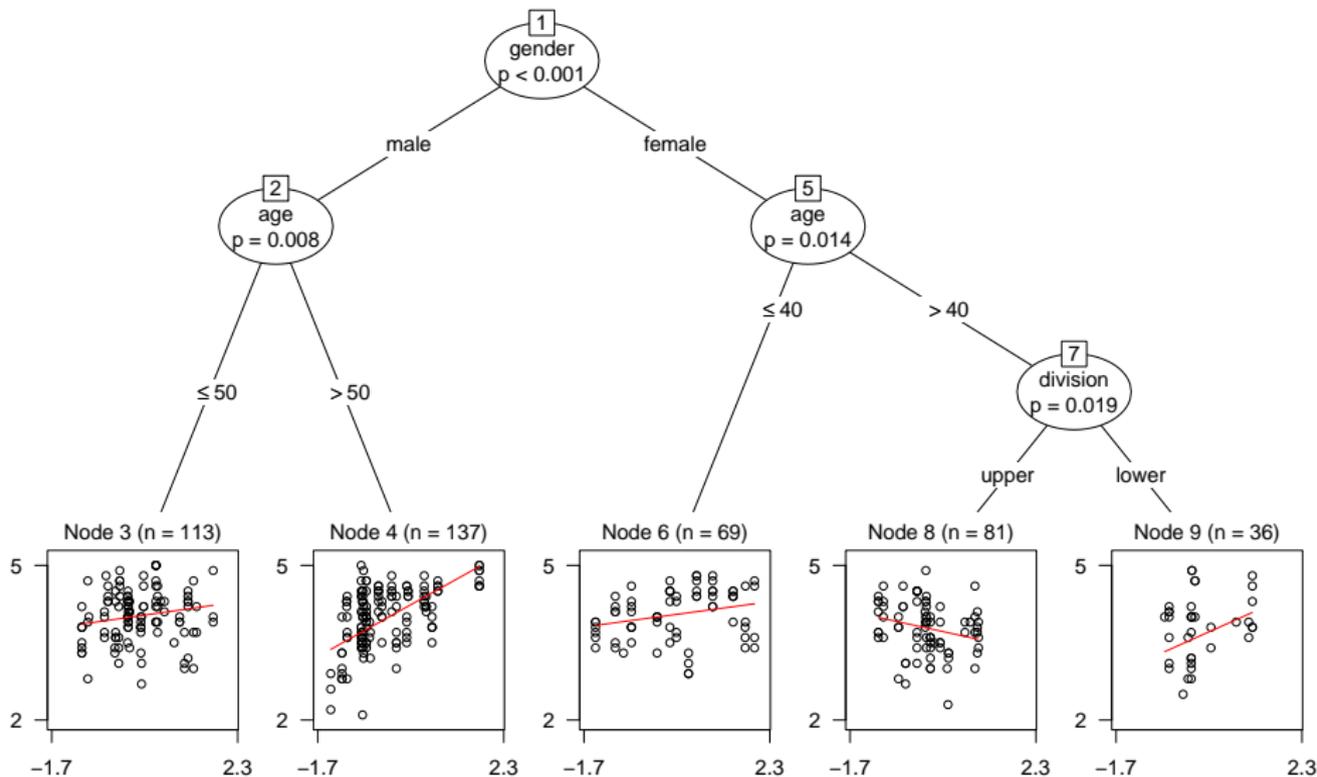
Hamermesh & Parker:

- Model with all factors (main effects).
- Improvement for separate models by gender.
- No association with age (linear or quadratic).

Here:

- Model for evaluation explained by beauty.
- Other variables as partitioning variables.
- Adaptive incorporation of correlations and interactions.

Beauty and teaching evaluation



Beauty and teaching evaluation

Recursive partitioning:

	(Const.)	Beauty
3	3.997	0.129
4	4.086	0.503
6	4.014	0.122
8	3.775	-0.198
9	3.590	0.403

Model comparison:

Model	R^2	Parameters
full sample	0.271	7
nested by gender	0.316	12
recursively partitioned	0.382	10 + 4

Software

All methods are implemented in the R system for statistical computing and graphics. Freely available under the GPL (General Public License) from the Comprehensive R Archive Network:

- Trees/recursive partytioning: **party**,
- Structural change inference: **strucchange**,

<http://www.R-project.org/>
<http://CRAN.R-project.org/>

Summary

Model-based recursive partitioning:

- Synthesis of classical parametric data models and algorithmic tree models.
- Based on modern class of parameter instability tests.
- Aims to minimize clearly defined objective function by greedy forward search.
- Can be applied general class of parametric models.
- Alternative to traditional means of model specification, especially for variables with unknown association.
- Object-oriented implementation freely available: Extension for new models requires some coding but not too extensive if interfaced model is well designed.